

On a characterization of the uniform repartition

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An important role in the theory of the hi-square criterion is played by the following fact: if x_1, x_2, \dots, x_n are independent random variables with Gauss distribution $N(0, \delta^2)$, then the distribution of the central statistic hi-square $\chi^2 = \sum_{i=1}^n (x_i + a_i)^2$ depends on a_1, a_2, \dots, a_n only by mean of the parameter $\sum_{i=1}^n a_i^2$. In the paper [1] one proves that this property is characteristic for the normal distribution of probability. The aim of this paper is to give a characterization of the uniform distribution of probability by mean of the hi-square statistic.

Theorem 1 *Let x_1, x_2, \dots, x_n independent and equally distributed random variables, where $n \geq 2$, then the necessary and sufficient condition that the statistic distribution $\chi^2 = \sum_{i=1}^n (x_i + a_i)^2$ depend on $\sum_{i=1}^n a_i^2$ with $a_i \in \mathbb{R}$ is that x_i be uniformly distributed.*

Proof. We define the function:

$$\psi(a) = Ee^{-(x_i+a)^2}. \quad (1)$$

It is obvious that $\psi(a) > 0$ and ψ is derivable in every $a \in \mathbb{R}$.

Using the conditions of the theorem we have

$$Ee^{-\sum_{i=1}^n (x_i+a_i)^2} = \prod_{i=1}^n Ee^{-(x_i+a_i)^2} = \prod_{i=1}^n \psi(a_i) = \Phi\left(\sum_{i=1}^n a_i\right). \quad (2)$$

Let $h(a) = \log \psi(a)$ and $H(a) = \log \Phi(a)$. From (2) we obtain:

$$\sum_{i=1}^n h(a_i) = H \left(\sum_{i=1}^n a_i \right). \quad (3)$$

If we differentiate twice the both sides of (3) by a_1 , then by a_2 , we obtain for every a_1, a_2, \dots, a_n :

$$H'' \left(\sum_{i=1}^n a_i \right) = 0. \quad (4)$$

In this way

$$H(a) = C_1 a + C_2. \quad (5)$$

From (1) and (3) we obtain:

$$\psi(a) = \int e^{-(x+a)^2} dF(x) = e^{C_1 a + C_3} \quad (6)$$

where $F(x) = P(x_i < x)$.

In the following step we consider the substitution:

$$e^{-x^2} dF(x) = dG. \quad (7)$$

In this case (6) can be written in the form:

$$\int e^{-2ax} dG(x) = e^{C_4 a + C_5}. \quad (8)$$

It follows, using the uniqueness theorem for the Laplace transformation, that $dG = C_5 \Delta(x - C_6)$ for every C_5 and C_6 , where Δ is the Dirac function. Using again relation (7), it follows that F is the distribution function of the uniform random variable.

The sufficiency can be proved by a straightforward verification.

References

- [1] Cagan A.M., Salcevski O.B. *Haracterizacija normalinovozaema, Svoisvom tetralinovo hi-cvadrat raspradelenia*, Litovski matem. Sbornic 1(1967) 57-58.

- [2] Cagan A.M., Salcevski O.B. *Dopustnii otenoc naimensti cvaderatov ischincitoenve cboistrov, normalinovo zacona*, Matematicheskie zematchi 6I(1969) 81-89.
- [3] Cagan A.M., Zingher A.A. *Sample mean as an estimator of location parameter. Case of non quadratic lass function* Sankhga Ser.A33(1971) 351-358.

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